

Breast-specific gamma camera imaging with ^{99m}Tc -MIBI has better diagnostic performance than magnetic resonance imaging in breast cancer patients: A meta-analysis

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Keywords: Specific gamma
imaging, -Magnetic resonance
imaging, -Breast cancer,
-Meta-analysis

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Received:

19 November 2016

Accepted revised:

23 December 2016

Abstract

Objective: This study aimed to evaluate the diagnostic role of breast-specific gamma camera imaging (BSGI) with technetium-99m-methoxy isobutyl isonitrile (^{99m}Tc -MIBI) and magnetic resonance imaging (MRI) in patients with breast cancer through a meta-analysis. **Subjects and Methods:** Three reviewers searched articles published in medical journals before June 2016 in MEDLINE, EMBASE and Springer Databases; the references listed in original articles were also retrieved. We used the quality assessment of diagnostic accuracy studies (QUADAS) tool to assess the quality of the included studies. Heterogeneity, pooled sensitivity and specificity, positive likelihood ratio, negative likelihood ratio, diagnostic odds ratio (DOR) and summary receiver operating characteristic (SROC) curves were calculated by Meta-DiSc software to estimate the diagnostic performance of BSGI and MRI. Ten studies with 517 patients were included after meeting the inclusion criteria. We did a subgroup analysis of the same data type. **Results:** The pooled sensitivities of BSGI and MRI were: 0.84 (95% CI, 0.79-0.88) and 0.89 (95% CI, 0.84-0.92) respectively, and the pooled specificities of BSGI and MRI were: 0.82 (95% CI, 0.74-0.88) and 0.39 (95% CI, 0.30-0.49) respectively. The areas under the SROC curve of BSGI and MRI were 0.93 and 0.72 respectively. **Conclusion:** The results of our meta-analysis indicated that compared with MRI, BSGI has similar sensitivity, higher specificity, better diagnostic performance, and can be widely used in clinical practice.

Hell J Nucl Med 2017; 20(1): 26-35

Epub ahead of print: 20 March 2017

Published online: 20 April 2017

Introduction

Breast cancer is the most frequently diagnosed cancer and the leading cause of cancer death in females, worldwide [1]. Mammography is a common method for screening breast cancer due to its high sensitivity, and ultrasound of the breast has been approved as an adjunct to mammography for women with dense breasts [2, 3]. However, some small and subtle breast lesions cannot be detected by mammography or ultrasound. For these lesions, magnetic resonance imaging (MRI) has a comparatively higher sensitivity [4]. Especially, contrast-enhanced MRI, can better depict the location and size of lesions and evaluate lymph node metastatic lesions not only on cross-section morphology but also on functional characteristics [5].

In recent years, as a non-invasive imaging modality and a kind of functional breast imaging technique, breast-specific gamma imaging (BSGI) with intravenous injection of technetium-99m-methoxy isobutyl isonitrile (MIBI) is increasingly used for diagnosis of breast cancer in clinical practice while its sensitivity is not influenced by breast density [6]. Breast-specific gamma imaging, also called molecular breast imaging, is a nuclear medicine breast imaging technique that uses a high resolution, small field-of-view breast-specific gamma camera [28]. Its radiotracer ^{99m}Tc -MIBI diffuses from the blood and mostly gathers in mitochondria. Uptake and retention of ^{99m}Tc -MIBI can reveal the level of tissue metabolism [30]. Compared with single photon emission tomography (SPET), BSGI offers superior intrinsic spatial resolution and can obtain standard mammographic views (craniocaudal [CC] and mediolateral oblique [MLO]) [7, 8]. Sensitivities and specificities of BSGI for the detection of breast cancer have been previously reported [6, 9-13].

Although there are many separate studies on BSGI and MRI, head-to-head comparison between these techniques has been rarely reported and includes only a few cases [21, 23-26]. Therefore, we conducted this meta-analysis to evaluate the diagnostic value of BSGI and MRI in the newly diagnosed or in the postoperative breast cancer patients.

Methods

Literature search strategy

We identified the published studies of the role of BSGI and MRI in the diagnosis of breast cancer in original articles by systematic searches of PubMed (including MEDLINE compiled by the United States National Library of Medicine, Bethesda, Maryland, USA) by using medical subject heading (MeSH) terms explicitly, Embase (Elsevier, Amsterdam, Netherlands), and Springer (Springer Science+Business Media, Berlin, Germany). The investigation was limited to studies conducted before June 2016. The keywords used were "Magnetic resonance imaging" or "MRI", "BSGI" or "Breast-specific gamma imaging" or "gamma imaging" and "Breast neoplasms or breast cancer or breast carcinoma".

Inclusion and exclusion criteria

Three reviewers independently screened titles and abstracts to select potentially eligible articles and manually searched reference lists of retrieved studies. Then we got the full text of the articles and determined their eligibility. Articles were included if they met the following criteria: a) Patients should have been examined both by BSGI and MRI before surgery; b) At least 10 patients were included in each study; c) Studies needed sufficient information for the calculation of sensitivity, specificity, and diagnostic accuracy; d) Reference standard should be completed, such as pathological examination or clinical and imaging follow-up; e) The images of BSGI and MRI should be visually interpreted by 2 experienced specialists. Exclusion criteria were as follows: a) Papers were reviews, conference abstracts, case report, letters, editorials or comments; b) Non-human studies; c) Non-English articles; d) Studies with duplicated data; e) Unpublished articles.

Quality assessment of the studies

Two authors independently assessed the risk of bias using the quality assessment of diagnostic accuracy studies (QUADAS) tool, which is a quality assessment tool specifically developed for systematic reviews of diagnostic accuracy studies. Fourteen items were included in the tool, covered the patient spectrum, reference standard, disease progression bias, verification bias, review bias, clinical review bias, incorporation bias, test execution, study withdrawals, and indeterminate results [14]. According to the guidelines, each of the items was scored to clarify studies as having high (yes), low (no) or unclear quality (unclear). The BSGI and MRI tests of each study were evaluated separately. The number of items evaluated as "yes" were regarded as QUADAS score of each study.

Data extraction

Two reviewers extracted data independently using a predefined data extraction form and any disagreements were resolved by discussion. The data extracted included the first author, study characteristics (i.e., country, year, and study design), patient characteristics (i.e., mean age, sample

size, and patient/lesion based), subgroup information (histologic and molecular classification), reference standard, state before examination (confirm/suspicious), injected dose of ^{99m}Tc -MIBI, image interpretation method of BSGI and MRI, as well as scan sequence of MRI (plain or contrast, type of functional imaging). True positive (TP), false positive (FP), false negative (FN) and true negative (TN) were extracted from each article.

Data synthesis and statistical analysis

We constructed a 2x2 contingency table for each study with results of TP, FP, FN, TN. By using these tables, pooled sensitivity and specificity, with 95% confidence intervals (CI), of two index tests of included studies were calculated, to compare the diagnostic accuracy of the two imaging modalities. A value of 1/2 was added to all empty cells or of studies to avoid potential problems in the calculation of summary estimates. Five studies only took pathology-proven patients into consideration. Therefore, the specificity of these studies was not reported and the FP and TN rates were recorded as 0.5. Positive likelihood ratio (LR^+), negative likelihood ratio (LR^-), diagnostic odds ratio (DOR) and summary receiver operating characteristic (SROC) curves were also calculated. The area under the curve (AUC) of the SROC was also used as an indicator of the performance of BSGI and MRI.

We used likelihood ratio I^2 index and χ^2 test (Chi-squared test or Q test) to assess heterogeneity. I^2 index is a measure of the percentage of total variation across studies due to heterogeneity beyond chance; a value over 50% indicates heterogeneity [15]. Regarding χ^2 test, when the p-value of a Q statistic was less than 0.05, it was considered as having apparent heterogeneity. We selected the random effects model (REM) when there was heterogeneity, otherwise we used the fixed effects model (FEM). A threshold effect was one of the main causes of heterogeneity when testing accuracy. The Spearman correlation coefficients for two index tests were calculated. If $P > 0.05$, there was no threshold effect [16]. We used Deeks' funnel plot to analyze the publication bias of all included studies.

All statistical analysis was completed using Meta-DiSc 1.4 (a free software to perform meta-analysis of studies of evaluations of diagnostic and screening tests) and Stata 12.0 (Stata Corporation, College Station, TX, USA).

Results

Study identification and selection

A total of 131 records were obtained in the initial search using the outlined search strategy. We found 16 articles considered potentially eligible and the full text was retrieved after reviewing title and abstract. Of the 16 full-text articles, 6 studies were excluded. Finally, we included 10 articles in our meta-analysis [17-26]. The flow diagram of the article selection process is presented in Figure 1.

Study characteristics

The characteristics of included studies are shown in Table 1.

The ten studies had a total of 590 subjects, including 517 patients, aged from 24 to 82. Four studies [18, 19, 23, 24] had a patients-based data type while eight [17, 20-26] had a lesion-based data type. Therefore, we included these eight studies of the same data type (lesion-based) to do a subgroup analysis of 393 lesions. The sample size of the studies ranged from 18 to 122 and the median sample size was 59. Five studies were from the USA, 4 from Korea and 1 from Austria. Among the 10 included publications, 7 were retrospective while 3 were prospective. Eight studies [19-26] reported histologic classification and 3 [19, 22, 24] reported molecular typing. All studies chose biopsy or postoperative pathology as the reference standard. Some studies have evaluated diagnostic value of mammography or ultrasound in patients with breast cancer, but we only extracted BSGI and MRI data. Among 10 included studies, 5 studies [22-26] lacked data of specificity because the study subjects initially selected were patients diagnosed with breast cancer by gold standard. Finally, only 3 studies [17, 20, 21] were available in the calculation of the pooled specificity according to lesions. Other details including imaging protocols and the diagnostic performance of each study are shown in Table 2.

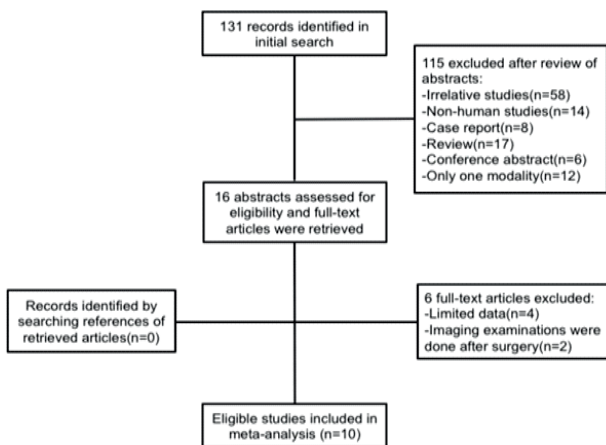


Figure 1. Flow chart of study selection process for meta-analysis.

Quality assessment

We used QUADAS tool to assess the studies included and the results of BSGI and MRI are reported separately. The QUADAS scores of each study are shown in Table 1. The quality assessment items and the likelihood of bias are summarized in Figure 2 and Figure 3. The reasons for unsatisfactory accordance of the representative spectrum (40% [17, 18, 21, 26]) were that half of the studies selected pathologically confirmed patients as subjects and that most studies were done in local institutions, resulting in a narrow spectrum of patients. The item of reference standard blinded was mostly reported unclearly, because most articles failed to provide specific execution of gold standard. On the whole, the qualities of the included studies were satisfactory and qualified.

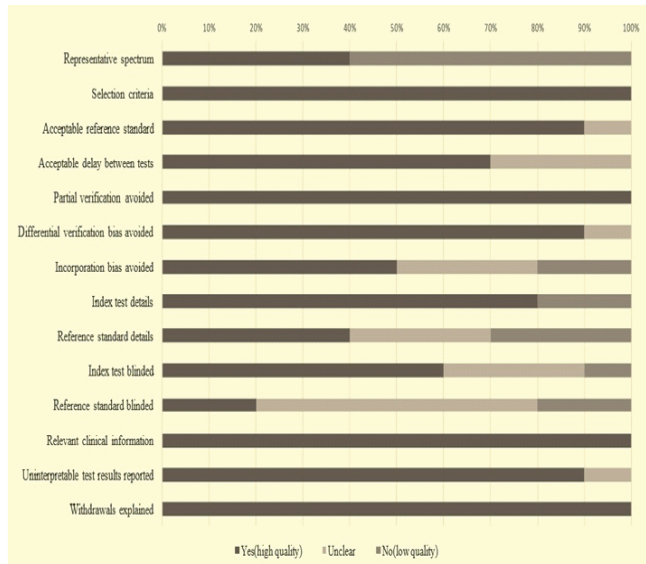


Figure 2. Methodological quality of included studies for BSGI according to the Quality Assessment of Diagnostic Accuracy Studies (QUADAS) tool.

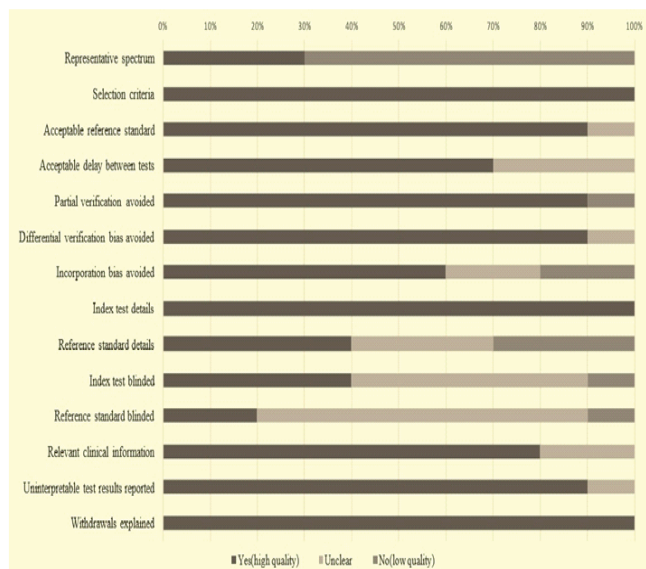


Figure 3. Methodological quality of included studies for MRI according to the Quality Assessment of Diagnostic Accuracy Studies (QUADAS) tool.

Heterogeneity test and publication bias

The heterogeneity test indicated statistical heterogeneity for two imaging modalities among studies, so we chose random effects model to calculate the pooled estimates. The statistical heterogeneity of BSGI and MRI are displayed in Figure 4 (sensitivity, Q value=9.1, I²=22.3%; specificity, Q value=8.30, I²=75.9%) and Figure 5 (sensitivity, Q value=9.42, I²=25.7%; specificity, Q value=4.12, I²=51.4%). The I² index can assess not only heterogeneity but also the degree of that heterogeneity as low (25%), moderate (50%), and high (75%) [26]. The Spearman correlation coefficient and P value of BSGI and MRI were 0.082 (P=0.846, P>0.05) and 0.709 (P=0.049, P<0.05), respectively. The result of MRI showed the existence of threshold effects in this meta-analysis. Deeks' Funnel plots of the included studies are presented in Figure 6, which showed insignificant publication bias (P=0.16, 0.61, respectively).

Table 1. Characteristics of the 10 studies included in the meta-analysis.

Author	Country	Year	Study design	Sample size	Patients/ Lesions	Median age (year)	Histologic subtype	Molecular subtype	Reference standard	QDAS score
Meissnitzer et al ¹⁷	Austria	2015	Pro	92	Lesions	>50	Benign(25) DCIS(3) IDC(53) Others(11)	NR	Biopsy	11(BSGI)
Johnson et al ¹⁸	USA	2014	Retro	75	Patients	NR	NR	NR	Biopsy	8(BSGI) 6(MRI)
Lee et al ¹⁹	Korea	2014	Retro	122	Patients	45.9(29-71)	IDC(115) Others(7)	ER+(65) PR+(49) HER2+(38)	Postoperative pathology	12(BSGI) 11(MRI)
Kim et al ²⁰	Korea	2012	Retro	97	Lesions	44.1	Benign(71) DCIS(6) IDC(16) Others(4)	NR	Biopsy	9(BSGI) 9(MRI)
Brem et al ²¹	USA	2007	Pro	33	Lesions	53(33-70)	Benign(24) DCIS(4) IDC(3) ILC(1) Others(1)	NR	Biopsy or follow-up	10(BSGI) 10(MRI)
Yoon et al ²²	Korea	2015	Retro	68	Lesions	49.5(24-73)	DCIS(58) IDC(10)	ER+(49) PR+(53) HER2+(26)	Biopsy and postoperative pathology	13(BSGI) 10(MRI)
Kim et al ²³	Korea	2014	Retro	35	Patients & lesions	48(26-69)	DCIS(35)	NR	Biopsy and postoperative pathology	10(BSGI) 10(MRI)
Keto et al ²⁴	USA	2012	Pro	18	Patients & lesions	51	DCIS(18)	ER+(16) PR+(12)	NR	8(BSGI) 8(MRI)
Brem et al ²⁵	USA	2009	Retro	28	Lesions	62.8(46-82)	ILC(28)	NR	Biopsy and postoperative pathology	9(BSGI) 8(MRI)
Brem et al ²⁶	USA	2007	Retro	22	Lesions	55(34-76)	DCIS(22)	NR	Biopsy and postoperative pathology	10(BSGI) 9(MRI)

(Pro, prospective; Retro, retrospective; NR, not reported; DCIS, ductal carcinoma in situ; IDC, invasive ductal carcinoma ILC, invasive lobular carcinoma)

Table 2. Imaging protocols and diagnostic performance of two image modalities.

Author	Country	Year	Image	State before examination	Image interpretation method	Injected dose of ^{99m} Tc-MIBI (MBq)	MRI sequence and functional MRI	Sensitivity	Specificity
Meissnitzer et al ¹⁷	Austria	2015	BSGI	Suspicious	Visual and semi-quantitative	740-1110		0.90	0.56
			MRI	>>	Visual		DCE-MRI, fat-suppressed T2, DWI	0.88	0.40
Johnson et al ¹⁸	USA	2014	BSGI	>>	>>	925-1110		0.92	0.73
			MRI	>>	NR		Unclear	0.89	0.54
Lee et al ¹⁹	Korea	2014	BSGI	>>	Visual	925-1110		0.740	0.722
			MRI	>>	NR		DCE-MRI	0.817	0.722
Kim et al ²⁰	Korea	2012	BSGI	Confirmed	Visual	740		0.888	0.901
			MRI	>>	>>		DCE-MRI, fat-suppressed T2, DWI	0.923	0.394
Brem et al ²¹	USA	2007	BSGI	Suspicious	>>	925-1110		0.89	0.71
			MRI	>>	>>		DCE-MRI	1.00	0.25
Yoon et al ²²	Korea	2015	BSGI	Confirmed	Visual and quantitative	555-925		0.83	NR
			MRI	>>	NR		DCE-MRI, fat-suppressed T2, DWI, ADC-MRI	0.94	NR
Kim et al ²³	Korea	2014	BGI	>>	Visual	925-1110		0.686	NR
			MRI	>>	>>		DCE-MRI	0.914	NR
Keto et al ²⁴	USA	2012	BSGI	>>	>>	925		0.89	NR
			MRI	>>	>>		E-MRI, MIP	0.94	NR
Brem et al ²⁵	USA	2009	BGI	>>	>>	740-925		0.93	NR
			MRI	>>	>>		Plain+ E-MRI, 3D volumetric imaging	0.83	NR
Brem et al ²⁶	USA	2007	BSGI	>>	>>	925-1110		0.91	
			MRI	Confirmed	>>		E-MRI, fat-suppressed T2, 3D volumetric dynamic imaging	0.88	NR

(DCE-MRI, dynamic contrast-enhanced MRI; E-MRI, enhanced-MRI; DWI, diffusion-weighted imaging; ADC-MRI, apparent diffusion coefficient MRI; MIP, maximum intensity projection; NR, not reported)

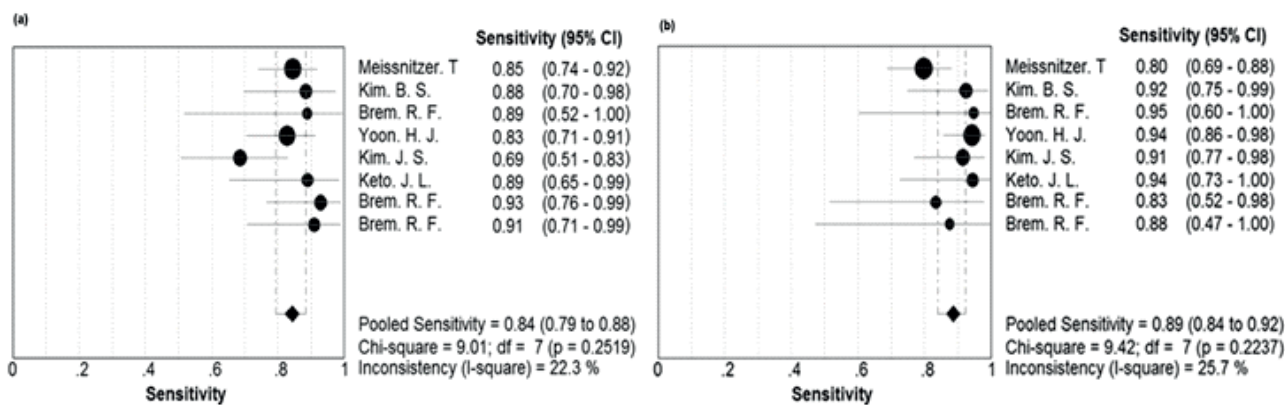


Figure 4. The pooled sensitivity of BSGI (a) and MRI (b). The size of the square plotting reflects the study weight. Horizontal lines are the 95% confidence intervals.

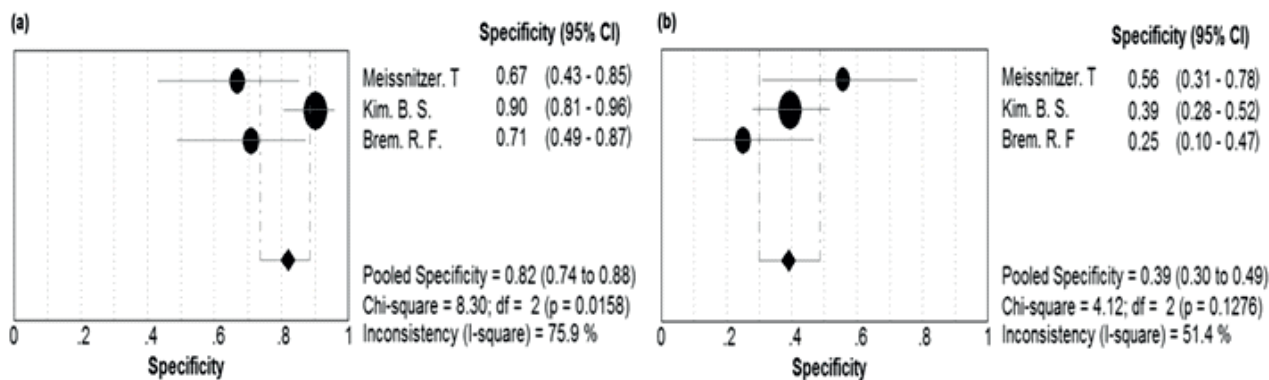


Figure 5. The pooled specificity of BSGI (a) and MRI (b). The size of the square plotting reflects the study weight. Horizontal lines are the 95% confidence intervals.

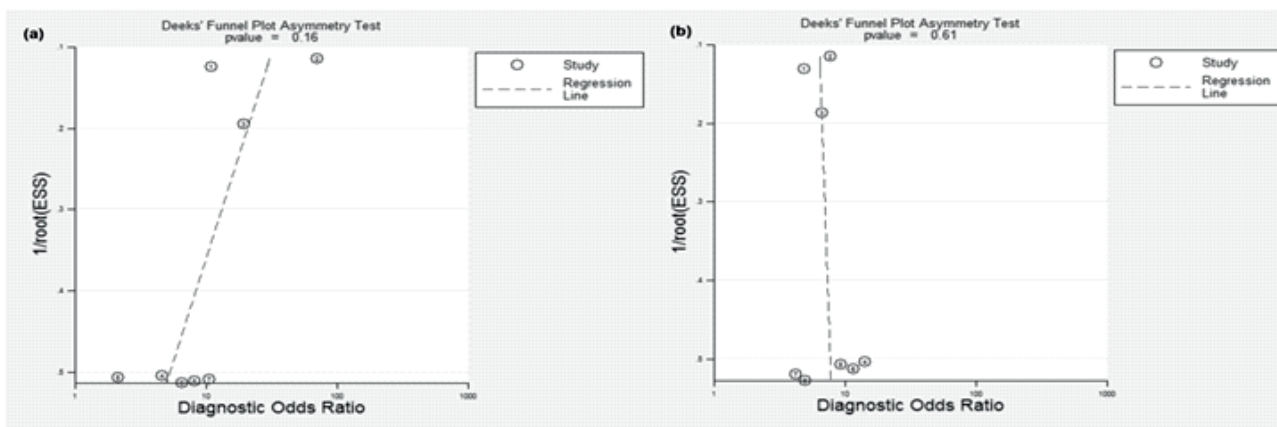


Figure 6. (a) Publication bias of BSGI using Deek's funnel plot. (b) Publication bias of MRI using Deek's funnel plot.

Diagnostic accuracy of BSGI and MRI

The pooled sensitivities of BSGI and MRI were 0.84 (95% CI, 0.79-0.88) and 0.89 (95% CI, 0.84-0.92), respectively, and the pooled specificities of BSGI and MRI were 0.82 (95% CI, 0.74-0.88) and 0.39 (95% CI, 0.30-0.49), respectively (Figure 4 and Figure 5). The results indicated that the sensitivity of MRI was slightly higher than BSGI, but the specificity of BSGI was significantly higher than MRI. For BSGI, positive likelihood ratio (LR⁺), negative likelihood ratio (LR⁻) and diagnostic odds ratio (DOR) were 3.15 (95% CI, 1.92-5.16), 0.21 (95% CI, 0.13-0.34) and 17.56 (95% CI, 8.32-37.09), respectively, while for MRI,

they were 1.46 (95% CI, 1.24-1.71), 0.29 (95% CI, 0.18-0.47) and 6.34 (95% CI, 2.92-13.75), respectively (Figure 7 and Figure 8). Summary received operating characteristics curves of BSGI and MRI were presented in Figure 9, the AUC values of which were 0.93 and 0.72, respectively, indicating a higher diagnostic accuracy of BSGI. The above results of this meta-analysis are presented in Table 3.

Discussion

Table 3. The overall diagnostic value of two image modalities.

Result	Sensitivity (95%CI)	Specificity (95%CI)	DOR	AUC	LR ⁺	LR ⁻
BSGI	0.84 (0.79 - 0.88)	0.82 (0.74 - 0.88)	17.56 (8.32 - 37.09)	0.93	3.15 (1.92 - 5.16)	0.21 (0.13 - 0.34)
MRI	0.89 (0.84 - 0.92)	0.39 (0.30 - 0.49)	6.34 (2.92 - 13.75)	0.72	1.46 (1.24 - 1.71)	0.29 (0.18 - 0.47)

(CI, confidence interval; DOR, diagnostic odds ratio; AUC, area under the curve; LR⁺, positive likelihood ratio; LR⁻, negative likelihood ratio)

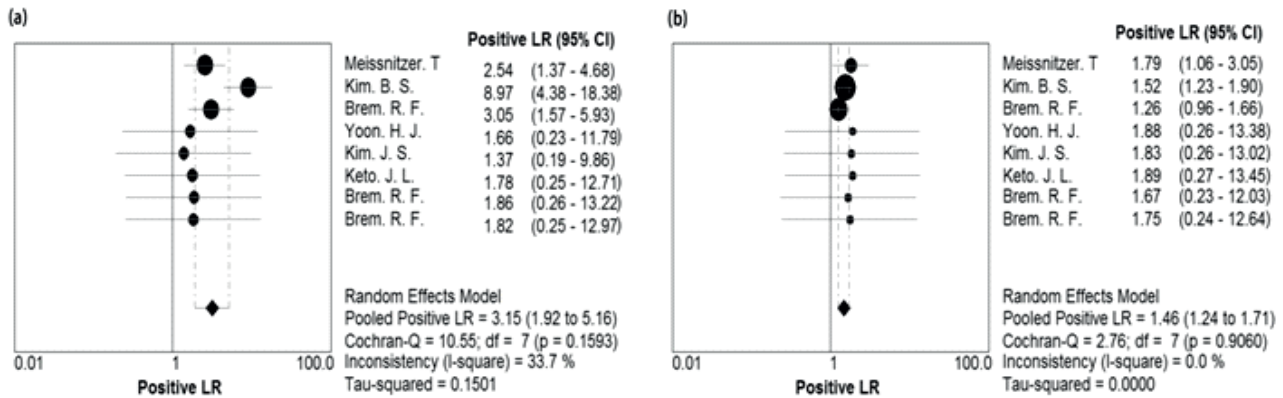


Figure 7. The pooled positive likelihood ratio of BSGI (a) and MRI (b). The size of the square plotting reflects the study weight. Horizontal lines are the 95% confidence intervals.

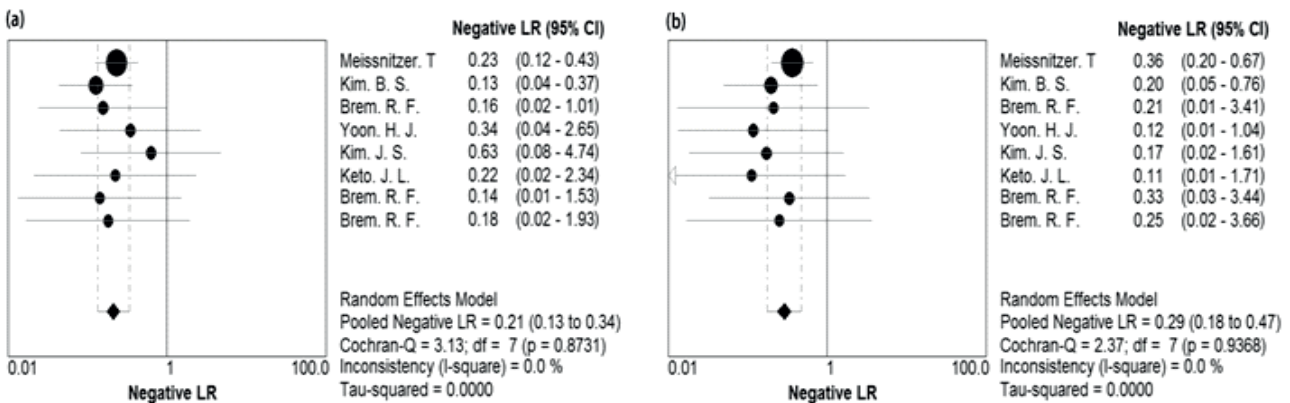


Figure 8. The pooled negative likelihood ratio of BSGI (a) and MRI (b). The size of the square plotting reflects the study weight. Horizontal lines are the 95% confidence intervals.

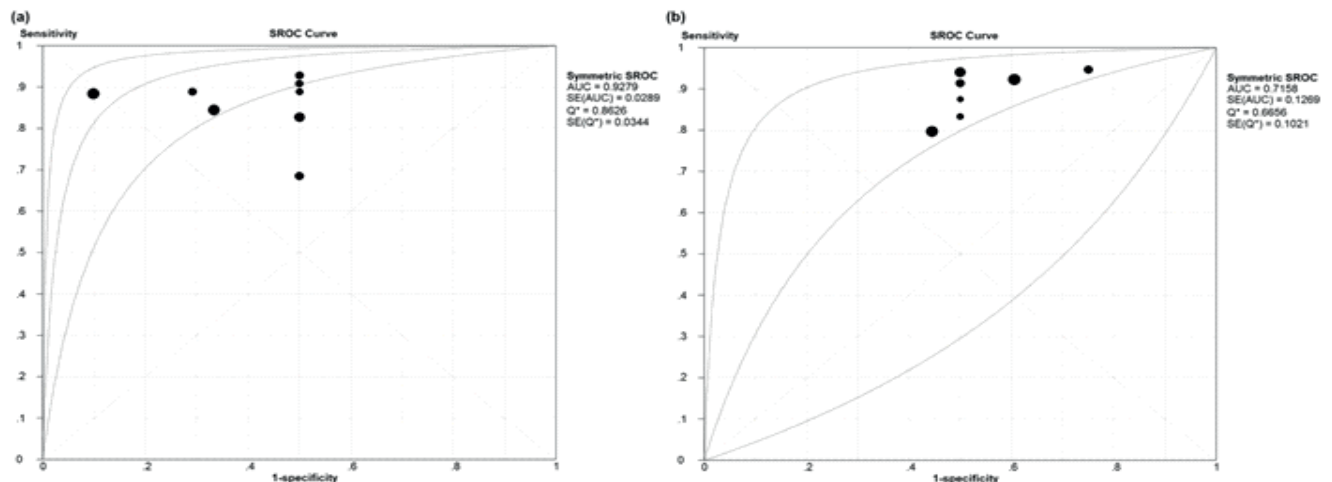


Figure 9. The pooled negative likelihood ratio of BSGI (a) and MRI (b). The size of the square plotting reflects the study weight. Horizontal lines are the 95% confidence intervals.

For diagnosis of breast cancer, Breast specific gamma imaging, as a new functional imaging modality, is increasingly used in clinical practice, while MRI, especially enhanced-MRI, is regarded as a routine method to make up for the limitations of mammography and ultrasound. Previous meta-analyses have suggested that BSGI had a high diagnostic performance as an excellent adjunct modality to mammography or ultrasound for detecting breast cancer [28] and MRI of the breast has high sensitivity and lower specificity in the evaluation of breast lesions [29]. In our study, pooled estimates and SROC curves were obtained to compare BSGI with MRI in detecting breast lesions. To our knowledge, this is the first meta-analysis to compare these two imaging modalities in patients with breast cancer.

In this study, we systematically evaluated the diagnostic value of BSGI and MRI in the same cohort patients with breast cancer. The results of the meta-analysis showed that compared with MRI, BSGI had sensitivity on par with MRI (84% vs 89%) but a higher specificity (82% vs 39%) and diagnostic efficacy (AUC 0.93 vs 0.72), indicating an excellent diagnostic performance.

The overall LR⁺ of BSGI was 3.15, and LR⁻ was 0.21, which were slightly different with the previous meta-analysis of Yu et al. (2013) [28]. In their study, the sensitivity, specificity, LR⁺, LR⁻ and AUC value of BSGI were 95%, 80%, 4.63, 0.08 and 0.9552, respectively. The difference might have resulted from the characteristics of lesions in included studies, like histopathological type and molecular type of cancers, which might reduce the sensitivity. Brem et al. (2007, 2009) [25, 26] reported that the sensitivity of BSGI in detecting invasive lobular carcinoma (ILC) was 93% while for ductal carcinoma in situ (DCIS) was 91%. Five studies [22-26] in our meta-analysis selected patients with lesions of a single pathological type as study objects, which may have contributed to the difference. Histopathological subgroup analysis cannot be performed due to limitations of the number of studies and information insufficiency of each study.

In all included studies, BSGI used ^{99m}Tc-MIBI as radiotracer. As a non-specific tracer, ^{99m}Tc-MIBI diffuses from the blood into the cytoplasm, and 90% gathers in mitochondria. Uptake and retention of ^{99m}Tc-MIBI depend on angiogenesis and regional perfusion, plasma and mitochondrial membrane potentials and thus the level of tissue metabolism [30]. Thus compared with the surrounding normal tissue, there is a higher uptake of ^{99m}Tc-MIBI in the cancer cells [31]. But breast benign hyperplasia lesions like fibrocystic change and fibroadenoma can also cause FP results in BSGI. False negative results appeared in some poorly-differentiated DCIS [21, 22, 26]. Heudel et al. [32] also found that uptake of ^{99m}Tc-MIBI in the 11 cases with lobular and tubulolobular carcinomas was lower than that in the ductal types. Some small lesions, especially sub-centimeter ones, could also cause FN. In the report of Meissnitzer et al. (2015), the sensitivity of BSGI for lesions <1 cm declined significantly to 60% [17].

Our results of MRI in the meta-analysis were comparable to those of Meissnitzer et al. (2015) [17] (sensitivity: 89% vs 88%, specificity: 39% vs 40%). On the whole, our results were concordant with theirs, but the application of different MR sequences may influence the results. All included studies

reported the application of enhanced-MRI except one [18], whose results had no significant effect on the overall results according to a sensitivity analysis. Three studies used DWI sequence [17, 20, 22] (Table 2), which is effective in distinguishing malignant and benign breast lesions with analysis of quantitative apparent diffusion coefficient (ADC) values [33]. Other MR techniques can also play important roles in the diagnosis of breast cancer, like fat-suppressed imaging, contrast-enhanced dynamic, maximum intensity projection (MIP) and sensitivity-encoding sense [34-36]. However, some benign hyperplasia lesions, micro enhanced lesions during the menstrual phase and local inflammation or necrosis after surgery or radiotherapy can cause FP results in MRI.

In addition to high specificity, which can help reduce numbers of biopsies of benign lesions, BSGI has the other advantages in the diagnosis of breast cancer: firstly, since Chinese women usually have smaller and denser breasts [40], the images of BSGI are not influenced by density and size of breasts. It is also relatively sensitive for small lesions and occult foci. Secondly, compared with MRI, some contraindications can be avoided, such as gadolinium allergies, claustrophobia, pacemaker and other metal implants. Instead of a prone position of MRI, BSGI obtains images in an upright seated position so the patients will be more comfortable. Furthermore, mediolateral oblique (MLO) views, as well as craniocaudal (CC) views allow a direct comparison to mammogram images. Thirdly, BSGI is more economical and efficient than MRI. Compared with hundreds of images obtained with MRI, only four to ten images are obtained with BSGI, the interpretation time of which is shorter as well. Zhou et al. (2008) [37] reported the total costs of BSGI and MRI in their institutions were \$1,259, \$3,400, respectively; Johnson et al. (2015) [18] reported institutional charges for BSGI and MRI were \$850 and \$3,381, respectively. Furthermore, BSGI is superior to MRI in the assessment of the response to neoadjuvant chemotherapy [19, 38] and is valuable for the preoperative evaluation of the extent of lesions. It can also help in the selection of a target for biopsy in the case of more than one lesion [17].

However, in detecting smaller lesions, morphological imaging modalities like MRI are superior to BSGI owing to a higher spatial resolution [17]. Breast specific gamma imaging also has limitations in detecting axillary lymph nodes and delineating adjacent lesions. In addition to these, we should not overlook radiation issue of ^{99m}Tc-MIBI. The injected dose of ^{99m}Tc-MIBI ranged from 555-1110MBq in all included studies (Table 2). Approximately 740-1110MBq of ^{99m}Tc-MIBI are recommended, and the effective dose equivalent for BSGI study is about 6.29-9.44mSv, which is relatively high compared with 0.7-1.0mSv of a screening mammogram [20]. There is a trend towards the application of BSGI in a low dose of ^{99m}Tc-MIBI in the future. Hruska et al. (2012) [39] have tried two methods to reduce dose: collimators optimized and widened energy acceptance window. They found that administered doses of 296MBq ^{99m}Tc-MIBI with the applied count sensitivity improvements permitted the detection of small breast lesions in patients. Their findings suggested that further reductions in acquisition duration or administered dose may be available [39]. More studies are needed for its clinical application.

Some limitations in our meta-analysis should be mentioned. First, some relevant articles might have been omitted even though we tried our best to retrieve medical literature. Second, there is not enough work on BSGI, as is a relatively new modality and is not widely used in clinical practice; therefore, the number of included studies is small and more ideal data types are needed. Some causes to heterogeneity or bias can hardly be avoided, due to the limited number of studies. Third, several factors may contribute to the heterogeneity; for example, seven studies were retrospective, except for three prospective studies with small sample sizes; studies included varied in image protocols and interpretation, methods of reference standard and ability of radiologists; among ten studies, five were from the US, four from Korea and only one from Europe, and three studies had the same Author. Unfortunately, it is difficult for us to find the exact source of heterogeneity due to the limited information.

Further studies on the role of BSGI in breast cancer are needed in the future, from lesion screening to evaluation of response to chemotherapy. We also need a larger sample for subgroup analyses, including different pathological subtypes and molecular subtypes.

In conclusion, our meta-analysis indicates that compared with MRI, BSGI has similar sensitivity but higher specificity and diagnostic performance. Breast specific gamma imaging can be widely used in clinical practice not only as an adjunct modality to mammography and ultrasound in the early detection of breast cancer, but also as a valuable examination in the assessment of the response to neoadjuvant chemotherapy and in directing the operation, especially when MRI is unavailable.

Acknowledgment

This work was funded by the National Natural Science Foundation of China (Grants No. 81471708), and Shanghai Jiao Tong University Medical Engineering Cross research fund (No.YG2012MS13).

The authors declare that they have no conflicts of interest

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